



**THE ROLE OF LOCALISED, RECOMBINANT AND
EXOGENOUS TECHNOLOGICAL CHANGE
IN EUROPEAN REGIONS**

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The role of localised, recombinant and exogenous technological change in European regions

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Abstract

How do regions develop and evolve along their productive and technological path is a central question. Within an evolutionary perspective, a given region is likely to develop new technologies closer to its pre-existing specialization. We adopt the approach of Hidalgo et al. (2007) to map the regional European technology/knowledge space to investigate the pattern and the evolution of regional specialisation in the most innovative EU countries. These dynamics depend on the interaction of three factors: (i) localised technological change, (ii) endogenous processes of knowledge recombination, and (iii) exogenous technological paradigm shifts while accounting for spatial and technological spillovers. Our paper maps the technological trajectories of 198 EU regions over the period 1986-2010 by using data on 121 patent sectors at the NUTS2 level for the 11 most innovative European countries, plus Switzerland and Norway. The results show that regional technological specialization is mainly shaped by localised technological change and exogenous technological paradigm shifts, whereas recombinant innovation contributes to a lower extent and that these effects largely depend on the increasing, decreasing or stable regional dynamics.

Keywords: Technology/knowledge space, localised technology change, recombinant innovation, European regions, evolutionary economic geography, patent analysis, spatial ordered models.

Jel classification: C23, O14, O31, O33, O52, R11, R12

1. Introduction

The technological and productive specialization of regions has always been a relevant issue from both a theoretical and empirical viewpoint. While globalisation and the ICT revolution have radically transformed the geography of production, in contrast to some early claims about “the Death of Distance” and “the World is Flat” (Cairncross, 1997 and Friedman, 2005), they have also spurred the importance of regional specialisation as a relative advantage in an increasingly competitive global arena.

The question of how regions develop and evolve along their productive and technological path is central in many scientific fields from international economics to economic geography, from public policy to regional science. Within an Evolutionary Economic Geography perspective (Boschma and Frenken 2006, 2011), we know that a region is most likely to develop new technologies and new industries closer to its previous technological and productive specialization. In other words, the current production structure sets the spectrum of prospects and limits to potential changes and developments: diversification mainly occurs by either using or recombining regional pre-existing competencies and know-how.

Recent progresses in the literature introduce other important notions and measures to analyse and possibly predict economic diversification in countries and regions. Hidalgo (2021) offers a rich and updated review, based on two main streams: the literature on relatedness focused on the evolution of specialisation (Boschma, 2017) and the literature on complexity centred on economic growth patterns (Balland and Rigby, 2017).

We follow mainly the former avenue of research by building on the conceptual blocks laid by Hausmann and Klinger (2007) who conceived an original methodology to map the evolution of industrial specialisation at the world level. Hidalgo et al. (2007) apply this methodology at the country level by using trade data to connect pairs of activities in terms of relatedness and estimate the affinity between locations and activities. Many other studies have moved along this path by applying this methodology at the regional level by using different data, territorial units and geographical settings. According to these studies, regions tend to diversify into new industries (Neffke et al. 2011; Boschma et al., 2012; Essleztbichler 2015; Xiao et al., 2018, Guo et al., 2018), new technologies (Colombelli et al., 2014, Kogler et al., 2013; Rigby 2015) and new occupations (Farinha et al. 2019), related to their present set of skills and capabilities.

However, according to Boschma (2021) and Pinheiro et al (2021), relatedness is only a part of the story since economic industrial dynamics are often a combination of related and unrelated diversification. The former is relatively more frequent, but the latter is, nevertheless, important and most notably essential to avoid regional economic lock-in.

In our analysis, we try to embody these two views by proposing a novel approach, which revisits, within a encompassing framework, three established concepts of the literature on the economics of innovation and technological change: (i) the notion of “localised technological change” conceived by Atkinson and Stiglitz

(1969), (ii) the idea of “recombinant growth” developed by Weitzman (1998), and iii) the concept of “exogenous innovation” invented by Kalecki (1954). According to this encompassing framework, incremental innovations develop along the lines of past technological “related” trajectories by causing local changes in the shape of isoquants rather than global shifts in their position. At the same time, new “unrelated” technologies may emerge from the recombination of existing technological knowledge, skills and capabilities. Finally, exogenous innovations determined by R&D efforts developed elsewhere might influence the technological specialization of a region.

We move along this stream of literature by trailing the study of Kogler et al. (2017) on the European knowledge space. In particular, our main aim is to investigate the evolution of European regional specialisation over more than two decades, from the middle eighties up to the complete outbreak of the economic crisis. We conceive the technological dynamics of a region as the outcome of the interaction of two endogenous - internal to each region - processes, i.e. localised technological change and recombinant innovation, together with an exogenous one, i.e. the shifts of the overall technological frontier. In other words, we acknowledge that *cases*¹, are not islands and that the technological evolution of a region cannot be entirely described by endogenous processes. It depends also on the interaction of some external factors such as the exogenous paradigm shifts of the global technological frontier. Moreover, we consider that there might also be proximity effects among external factors due to the interaction across regions in either the geographic or the technological dimension (De Dominicis et al., 2013 and Paci et al., 2014).

The main original and innovative contributions of the paper are threefold. First, we attempt to discriminate between technological changes, which happen because of either endogenous (local) or exogenous (global) shifts. Secondly, we propose a new way to operationalise the concept of relatedness/unrelatedness through the concepts of degree and betweenness centrality indexes derived from social network analysis (SNA) applied to European Maximum Spanning Trees (MST). Thirdly, by applying the Correlated Random Effects (CRE) estimation approach (Wooldridge, 2005 and 2010) to ordered logit panel models, we can exploit both the within and the between variation in our data and thus provide a more reliable estimation of the effects on regional technological specialization exerted by the main explanatory variables considered in our analysis.

The analysis is based on patent data for 121 International Patent Classification (IPC) classes for 198 European NUTS2 regions observed over the period 1986-2010. The main results show that regional technological specialization is mainly shaped by localised technological change and exogenous technological paradigm shifts, whereas processes of recombinant innovation contribute to a lower extent. Moreover, results point out that, once accounting for spatial and technological spillovers and transition dynamics, it is the between variation which

¹ In this paper, we refer to a “case” as a shortcut for the combination of a specific technological sector within a given region.

plays a major role. When we split the sample to focus specifically on cases recording either increasing, stable, or decreasing specializations, previous results are confirmed and new evidence emerges. For stable cases, our exogenous variables exert their greater influence when a region has not developed any technological sector, followed by the case in which its technological comparative advantage records values around one. When technological specialization is increasing over time, the largest effect is found for the very specialised cases. Finally, when specialization decreases, the largest influence is recorded in the lowest possible state, showing the role of contiguous IPC classes in avoiding the total disappearance of a given class in the spectrum of the technological specialization of a region.

The paper is organised as follows. The next section connects this analysis to the established literature in the field, section three describes how the European technological space is mapped by applying SNA to MST, derived from technological interrelatedness matrices; some stylised facts of the European technological space are then presented, along with its regional evolution over time. The fourth section introduces the main explanatory variables, while the fifth section presents the estimation strategy. The empirical analysis results are presented and discussed in section six. Concluding remarks and future research agenda are in the final section.

2. Literature review

The prevalent model of technological change used in empirical analyses – the Knowledge Production Function (KPF, Griliches, 1979) – assumes that the most significant source of knowledge, besides human capital and skilled labour, is public and private R&D. This empirical model has been applied at different levels of economic systems: from the micro level of plants and firms to the macro level of industries, regions and nations. Nonetheless, the mechanical idea of knowledge creation with the indirect assumption of almost perfect plasticity of the innovative structure of a region – a linear production process where R&D is the input and innovation is the output – is not entirely satisfactory. Even Griliches himself, in his conclusion, acknowledges that “We need more research on ... how to conceptualise and estimate the technological distance between firms and industries and the associated notions of externalities and spillovers in research” (Griliches, 1979, p. 43).

We, therefore, go beyond the classical KPF approach to rediscover three established theoretical contributions: localised technological change (Atkinson and Stiglitz, 1969), recombinant growth (Weitzman 1998) and exogenous innovation (Kalecki, 1954). Our basic analytical framework grafts these theoretical contributions and instruments them through the use of SNA applied to MSTs within the current literature of evolutionary economic geography (Boschma and Frenken, 2011). According to this encompassing framework, incremental innovations develop along the lines of past technological “related” trajectories by causing local changes in the shape of isoquants rather than global shifts in their position. At the same time, new “unrelated” technologies may emerge either from the recombination of locally

available existing skills and capabilities or the adoption of exogenously created technological knowledge.

A practical solution for the operationalization of these two propositions is borrowed by Hidalgo et al. (2007), who describe technology as based on the current competencies of countries and regions and centre their approach on the notion of “technological proximity” and the representation of the technological frontier through a MST. Hidalgo and Hausmann (2008) use a persuasive metaphor for their methodology: products are trees, and forests compose the economic structure of countries; firms are, instead, monkeys that live on different trees and exploit those products. Growth and specialization dynamics can be described as the movement of firms from a poorer part of the forest to more prosperous parts of the forest, where trees bear more fruits and develop faster. This metaphor is essential to appreciate the concept of interrelatedness: “if this forest is heterogeneous, with some dense areas and other more-deserted ones, and if monkeys can jump only limited distances, then monkeys may be unable to move through the forest.” (Shaw, 2010, p. 8).

Consequently, the composition and the relative density of a forest, that is, the economic structure of a country/region, is crucial in determining the orientation and the pace of development of countries/regions in the short and long run. Hidalgo et al. (2007) employ this method to show that rich countries specialize in more densely connected parts of the product space while developing countries mainly develop products in the more peripheral and isolated areas. As a result, rich countries have more opportunities to sustain economic growth in the end, thanks to a fruitful process of structural change and to their diffused absorptive capacities. In other words, the probability of success for a region entering a new economic activity depends on geographical closeness and the cognitive and technological proximity between the new activity and a region’s prior activities. Furthermore, Boschma (2005) asserts that relatedness and proximity are crucial in favouring changes, not only in the geographic and technological space. Other dimensions may prove essential, such as social, organisational and institutional proximity as potentially favouring factors for knowledge spillovers.

These concepts have been mostly applied in the literature at the regional level since technological knowledge has a tacit nature and can be strongly associated with local capabilities, institutional settings and social capital (Lawson, 1999; Breschi, 2000, Maggioni, 2002; Greunz, 2003; Moreno et al., 2005). Regions may, therefore, accumulate specific competencies and intangible assets, which provide spatially and cognitively bounded learning opportunities for local firms. In other words, this seems the ideal geographical setting for analysing the evolution of the regional technological system as the result of a set of local factors. This belief was also at the base of some studies which apply the metaphor of trees and monkeys to the impact of technological relatedness on regions within a single country².

² Earlier works focus on regional economic growth. Boschma and Iammarino (2009) and Boschma et al. (2012) show that relatedness is an essential component in raising the opportunities to grow in provinces in Italy and in Spain, respectively. Neffke et al. (2011) is more directly oriented to the issue

The first study by Boschma et al. (2015) investigates, thanks to US Patent and Trademark Office (USPTO) patents, the role of technological relatedness in pushing and orienteering technological change in 366 US cities (MSA) from 1981 to 2010. They find that technological relatedness may play a crucial role by increasing a new technology's entry probability and decreasing existing technology's exit probability. They use two different methods to build the relatedness indicator. The primary method follows the usual product space framework proposed by Hidalgo et al. (2007), where two technologies are considered related if they have a revealed technological advantage within the same US city. The second method, used to test the robustness of the results, is based on Hall et al. (2001) patent classification and a normalised co-occurrence analysis.

Kogler et al., (2013) have a similar objective and methodology. They base their measure of relatedness utilising co-classification information contained in patent documents. They show that over time, patents increasingly cluster within technology classes close to one another in the technology space. They also reveal considerable heterogeneity in measures of technological specialisation across US metropolitan areas. In general, smaller cities tend to display higher levels of knowledge relatedness, often because a small number of firms with a limited range of technological know-how controls the invention process. Larger cities generate more broadly dispersed knowledge across the US knowledge space.

Rigby (2015) studies the evolution of knowledge space again in the same sample of 366 MSA from 1975 to 2005. Technological relatedness is now based on patent citations, and it is given by the probability that a patent in class j will cite a patent in class i . The analysis shows that the average relatedness between US patents in thirty years has almost doubled since patents are increasingly concentrating in fewer technology classes, which are becoming more proximate (or related). Regarding the determinant of entries and exits of cities from patent classes, the expansion of the knowledge core depends on the proximity of new technological possibilities to the set of existing specialisations. Most interestingly, estimations show that other dimensions of proximity, other than the technological one, play a role: diversification is also influenced by the knowledge available in socially closer locations, where social proximity is measured in terms of co-inventors links.³

of industrial branching within regions. They study products entry and exit in 70 regions in Sweden by looking at employment data and measuring technological relatedness thanks to an original dataset on product co-occurrences in plants.

³ US Metropolitan areas are also at the center of Essletzbichler (2015) analysis, even though the relatedness measure is based on input-output linkages between industries rather than patent or products co-occurrences. Nonetheless, results confirm that technological relatedness is positively related to previous industry portfolio membership and industry entry and negatively related to industry exit. The latest contribution on the US case is by Farinha et al. (2019) who unpack relatedness to distinguish between three mechanisms: complementarity (interdependent tasks), similarity (sharing similar skills) and local synergy (based on pure co-location). They assess their impact on the evolution of the occupational structure of 389 US Metropolitan Statistical Areas (MSA) for the period 2005–2016. They find, as expected, that new jobs are related to existing ones, while those ones disappearing are more unrelated to a city's jobs' portfolio. They find that all three relatedness dimensions matter, but local synergy shows the largest impact on entry and exit of jobs in US cities.

Another parallel avenue of research focused on the EU. This focus started with Kogler et al. (2017), who use patent co-classification data to measure the proximity between all pairs of IPC categories to map and track the evolution of knowledge space from 1981 to 2005 in 213 NUTS2 regions of EU15. They find that, as in the US, in Europe knowledge specialisation has increased significantly over time. They also show that entry, exit and selection processes over space and time are influenced not only by the proximity to the knowledge core of the region but also by knowledge spillovers from neighbouring regions.

Similarly, Xiao et al. (2018), by using employment data from the Orbis database across 173 European regions during the period 2004–2012, show that the probability of a new industry specialisation in a region is positively associated with its relatedness to the region's current industries. Moreover, they prove that the influence of relatedness on the probability of new industrial specialisations depends on the innovation capacity of a region: relatedness is more relevant in weaker regions in terms of innovation capacity. This result implies that more vulnerable regions are more inertial, whilst most innovative areas are more able to break their technological path, that is, their potential lock-in.

These pioneering works on relatedness were later enriched by widening the scope of analysis to complexity in EU regions, thanks to Rigby et al. (2021). They contribute by mapping the trajectories of EU city-regions in a smart specialisation space from 1981 to 2015. They use panel models to show that employment and GDP grow faster in cities that build capabilities in complex new technologies close to their existing knowledge cores while abandoning less complex, unrelated technologies.

Lately, along the same research avenue, Pinheiro et al. (2022) use data on industries and patents to analyse the diversification patterns of 283 regions in 32 European countries over the past 15 years. Only the most economically advanced regions could diversify into highly complex activities. These regions tend to focus on related high-complex activities, while lagging regions focus on related low-complex activities, creating a spatial inequality feedback loop. This pattern creates a wicked problem for innovation policy: the strategy needed to improve the innovativeness of the European knowledge system might disproportionately benefit already developed regions and foster disparities.

In this contribution, we consider the technological relatedness within regions controlling for the possibility that spillovers may come from both geographical and technological proximity. In addition, we contribute to the empirical literature on technological diversification in Europe by assessing how specialization at the regional level is driven by the two internal factors of localised technological change and recombinant innovation, as well as by exogenous shifts of the technological frontier. As described in section 4, all three factors are operationalized in a novel way as they are directly derived from the European knowledge space obtained by applying SNA derived indexes to the MST method as in Hidalgo et al. (2007).

3. The European knowledge space

To investigate the pattern and the evolution of the technological specialisation of the EU regions we start by depicting the most salient traits of the European knowledge space obtained using the MST method. To construct the MSTs, which refers to the period 1986-2010, we follow the approach proposed by Hidalgo et al. (2007), based on the notion of co-specialization.

MSTs are built by using data on the number of patent applications filed at the European Patent Office (EPO) classified by priority year and by inventor's region for 198 NUTS2 regions in Europe (EU13), belonging to the most innovative countries in Europe (see table A1 in the Appendix), recording 97% of total European patents in the period 1986-2010.

Since patenting activity at the regional level is quite erratic over time, we smooth the patent variable by computing five-year period averages. Therefore, our analysis is articulated in 5 five-year periods⁴.

To deal with the sectoral dimension of technological interrelatedness, we focus on 121 IPC classes at the second hierarchical level.

To describe the technological interrelatedness between IPC classes for EU13 regions, we start from the “innovation space”, a notion similar to the “product space” as defined by Hidalgo et al. (2007). The “innovation space” is, in principle, a connectivity matrix, which shows how closely interrelated is a sector, that is an IPC class, with another one. In our study, it is a 121x121 matrix whose rows and columns represent IPC classes and each off-diagonal cell represents two measures of technological connectivity between a given pair of IPC classes.

This matrix can be interpreted as a network, where each node is an IPC class and each link measures the relatedness between two IPC classes.

More specifically, the CO-Specialization (CO-SP) approach is implemented using the Hidalgo et al. (2007) MST method, which relies on conditional probabilities of a region being specialized in IPC class i given that the same region is also specialized in IPC class j . Regional specialization in a given IPC class is measured in terms of Revealed Technological Advantage (RTA). This is computed as the proportion of a region's patents (pat) in a given IPC class i in period t , divided by the proportion of European patents in the same IPC class in the same period. Formally:

$$RTA_{ir}^t = \frac{\frac{pat_{ir}^t}{\sum_i^I pat_{ir}^t}}{\frac{\sum_r^R pat_{ir}^t}{\sum_r^R \sum_i^I pat_{ir}^t}}$$

where $i = 1, \dots, I = 121$ IPC, $r = 1, \dots, R = 198$ regions, $t = 1, \dots, T = 5$ five-year periods.

⁴ Time intervals are as follows: 1986-90 (T1), 1991-95 (T2), 1996-2000 (T3), 2001-05 (T4), 2006-10 (T5).

Given the conditional probabilities $P(RTA_{ir}^t | RTA_{jr}^t)$ and $P(RTA_{jr}^t | RTA_{ir}^t)$ of a region being specialized at time t in IPC class i given that the same region is also specialized in IPC class j and vice versa, each element of the CO-SP connectivity matrix at time t , $co_sp_{ij}^t$ is equal to:

$$co_sp_{ij}^t = \min\{P(RTA_{ir}^t | RTA_{jr}^t), P(RTA_{jr}^t | RTA_{ir}^t)\}$$

Since the CO-SP matrix is very dense and links values are very heterogeneous, we decided to focus our analysis only on key technological relations underlying the European technological space. To do so, following Hidalgo et al. (2007), for each interval of time, we identified a European MST, whose nodes are 121 IPC classes and links include exclusively the most relevant technological interrelation between a couple of IPC classes.

The procedure to create the MSTs is iterative and starts with the identification of the maximum link value in each connectivity network. Once the maximum value has been selected, we establish a link between that couplet of IPC classes or nodes. Secondly, by focusing on the identified dyad, we search for a further node to be connected to that dyad in order to form a triad. The link is identified by searching for the maximum value of all links attached to one of the two nodes of the dyad. The procedure is iteratively repeated by adding nodes (i.e. IPC classes) and links until all IPC classes are included in the MST (which, by definition is a minimally dense network of N nodes and $N-1$ links). The final MST structure for the initial and the last time period (T1: 1986-1990 and T5: 2005-2010) is depicted in Figure 1.

FIGURE 1 HERE

By looking at the temporal evolution of the CO-SP MSTs from 1986 to 2010 it is evident that significant changes in the MST structure have occurred. Thus, looking for a standardized measure of correlations between different networks we resolved in using the Quadratic Assignment Procedure (QAP), to calculate the extent to which the pattern of links in one period is correlated with the pattern in another period. Standard correlation is not appropriate for dyadic data because such data are not independent of each other. QAP controls for the non-independence of the cases using several random permutations of rows and columns of the original matrix through a Monte Carlo procedure, thus it allows to rule out spurious correlations (Krackhardt, 1988). Table A2 in the Appendix shows relatively low level of association across time: all QAP correlation coefficients (in the range 0.17-0.25) are significant and indicate positive autocorrelation, whose strength tends to decrease as the lag length increases. The reduction in the correlation values may be interpreted as a sign of the incremental nature of technological change, as time passes the technological frontiers keeps modifying based on the previous one.

For robustness purposes, we also compute the European knowledge space by applying a method similar to the one proposed by Engelsman and Van Raan

(1992). According to this approach, “two technology classes are considered to be technologically related if they occur frequently together as technology classification codes on the same patent” (p. 6). Thus, each element of the CO-CLassification (CO-CL) connectivity matrix (c_{ij}) is computed for each period as the number of EU13 patents in which a given couplet of IPC classes i, j is jointly occurring. CO-SP adjacency matrix will be used to define regressors as in section 4.2.

4. Variables description

As stated in the introduction, the main aim of this study is to explain the evolution of the technological specialization of the European regions as a function of localized technological change (LTC), recombinant innovation (RI), and exogenous technological innovation (ETC), while accounting at the same time for technological and spatial spillovers. We carry out the empirical analysis along two perspectives: a “static” one, detected with the dependent variable named $SRTA_{ir}^t$, aimed at capturing the determinants of the state of the art of technological specialisation, and a “dynamic” one aimed at identifying the main drivers of changes and evolution in the innovation process, i.e. the factors leading to increasing or decreasing levels of specialisation. In the next sub-sections, we describe the dependent variable and then we focus on the main explanatory variables⁵.

4.1 The dependent variable

Since our measure of relative specialisation is given by RTA, as defined in section 3, which is a continuous variable, in a *static* perspective we define five possible ordered categories that we use to characterize patterns of technological specialisation of European regions and we name the variable $SRTA_{ir}^t$. Thus, similarly to Guevara et al. (2016), we discretize the original continuous variable RTA_{ir}^t to deal with the pronounced asymmetry of its distribution (asymmetry value is equal to 23.335 and kurtosis is equal to 1214.631). These five states are:

1. *Inactive* if $RTA_{ir}^t = 0$
2. *Very Unspecialised* if $0 < RTA_{ir}^t \leq 0.5$
3. *Unspecialised* if $0.5 < RTA_{ir}^t \leq 1$
4. *Specialised* if $1 < RTA_{ir}^t \leq 1.5$
5. *Very Specialised* if $RTA_{ir}^t > 1.5$

The distribution of the resulting dependent variable, $SRTA_{ir}^t$, describing the technological specialization of European regions, is depicted in Figure 2. Throughout the period analysed, *Inactive* cases counts for nearly 28% of the total,

⁵ A complete description of all variables, along with basic descriptive statistics, is reported in Table A3 in the Appendix.

Very Unspecialised cases are 23.5%, *Unspecialised* cases reach 19.1%, *Specialised* cases are 10.5%, *Very Specialised* cases are 18.9%.

FIGURE 2 and 3 HERE

During the period under consideration, the technological specialisation of regions records significant variation, thus we investigate the specialisation dynamics of European regions, by identifying three groups:

1. *Stable Specialization Dynamics*, if the value of $SRTA_{it}$ remains in the same state in both t and $t+1$.
2. *Increased Specialization Dynamics*, if the value of $SRTA_{it}$ records an upward shift in the states from t to $t+1$.
3. *Decreased Specialization Dynamics*, if the value of $SRTA_{it}$ records a downward shift in the states from t to $t+1$.

Figure 3 shows the distribution of the three groups according to this taxonomy. Interestingly, most cases can be classified as stable (51.9%), while the increased dynamic group (24.8) is slightly larger than the decreased dynamic one (23.3). These figures show a substantial stability of the technological specialisation of the European regions with a marginal but significant bias towards an increased specialization.

TABLE 1 HERE

Panel A of Table 1 reports a matrix with the absolute numbers for the three groups of cases: no change cases are on the main diagonal (stable specialization), forward changes in above-diagonal cells (increased dynamic) and backward changes in below-diagonal cells changes (decreased dynamic). Panel B of Table 1 presents data disaggregated by the number of forward or backward steps, whereas Panel C reports the percentage values of each of 5 states in each of the 3 groups. While the inactive and the very unspecialised cases mostly display stable dynamic, the very specialised cases witness an increased dynamic and the inactive and very unspecialised display a decreased dynamic. As shown in panel C, while stable dynamic show all 5 ordered states, for the other two groups, there are only 4 possible ordered states (i.e. from *very unspecialised* to *very specialized* state for the *increased* dynamic; from *inactive* to the *specialized* state for the *decreased* dynamic).

In synthesis, in the first part of the econometric analysis, we assess the role of localized technological change, recombinant innovation, and exogenous technological shift in explaining the outcomes of the dependent variable represented by the five ordered states of specialization; while in the second part, the analysis is carried out by 3 groups of cases, i.e. *stable*, *increased* and *decreased dynamics*.

4.2 The explanatory variables

As anticipated in the Introduction, technological specialization is the result of both related and unrelated diversification processes, as well as changes in the European

technological frontier, exogenous to any single region. In our analysis, we measure relatedness, unrelatedness and exogeneity by making use of the notion of “localised technological change” (Atkinson and Stiglitz, 1969), the concept of “recombinant growth” (Weitzman, 1998), and the notion of exogenous innovation (Kalecki, 1954) through the development of three original indexes based on the application of the principles of the SNA to MSTs. All explanatory variables are computed on the basis of the CO-SP MST described in Section 3. In what follows we provide a detailed description of the procedure used to construct each variable.

Localized Technological Change (LTC)

LTC is the variable implementing Atkinson and Stiglitz’s claim that “the different points on the [production possibilities] curve represent different processes of production and associated with each of these processes there will be certain technical knowledge specific to that technique. Indeed, both supporters and critics of the neoclassical theory seem to have missed one of the most important points of the activity analysis (Mrs Robinson’s blueprint) approach: that if one brings about a technological improvement in one of the blueprints this may have little or no effect on the other blueprints. If the effect of technological advance is to improve one technique of production but no other techniques of producing the same product, then the resulting change in the production function is represented by an outward movement at one point and not a general shift. In reality, we should expect that a given technical advance would give rise to some spillovers and that several techniques would be affected” (Atkinson and Stiglitz, 1969, p. 573).

To operationalise the direct “effects of technological advance”, we exploited the concept of adjacency, and the related degree centralization index, derived from SNA.

Given a set of nodes $\mathcal{N} = \{n_1, n_2, \dots, n_{121}\}$, there are several paths, with different lengths, connecting a given pair of nodes. The shortest path between two nodes i and j is named geodesic distance and is denoted as g_{ij} . If $g_{ij} = 1$, nodes are adjacent, indicating that there exists a direct link between them, otherwise if $g_{ij} > 1$, nodes are not directly linked and the number indicates the smallest length connecting them. Hence, to detect the direct, or local, technological effects we used the concept of adjacency of nodes. In this case, the local neighborhood is defined as $g_{ij} = 1$. Therefore, LTC is obtained by computing per each IPC class, each region, and time the summation of RTA of nodes directly adjacent to the MSTs. Formally:

$$LTC_{ir}^t = \sum_{j=1}^{N-1} RTA_{jr}^t | (g_{ij} = 1, MST^t)$$

In synthesis, LTC explains the technological specialization of a region, in a given IPC class, in terms of the technological specialization of the same region in related (adjacent) technologies.

Recombinant Innovation (RI)

RI is the variable implementing Weitzman's claims that "[In the knowledge production function approach] 'New ideas' are simply taken to be some exogenously determined function of 'research effort' in the spirit of a humdrum conventional relationship between inputs and outputs. Essentially, this approach represents a theory of knowledge production that tries to do an end run around describing the creative act that produces the new ideas. If new ideas are postulated to be a function of something – for example, research effort – then what is the nature of the functional relationship? Is the production of knowledge a process that can be modelled by analogy with fishing new ponds or discovering new oil reserves? It seems to me that something fundamentally different is involved here. When research effort is applied, new ideas arise out of existing ideas in some kind of cumulative interactive process that intuitively seems somewhat different from prospecting for petroleum. To me, the research process has at its centre a sort of pattern-fitting or combinatoric feel. The core of the analytical structure is a theory of innovation based on an analogy with the development of new cultivated varieties by an agricultural research station. 'Recombinant innovation' refers to the way that old ideas can be reconfigured in new ways to make new ideas" (Weitzman, 1998, p. 332-333).

In order to operationalise the concept of new - possibly complex - recombination of old knowledge we adapted, from SNA, the concept of *betweenness centrality*, i.e. an analytical measure of the strategic role played by a node when lying between the geodesic paths connecting other nodes in the network (Freeman, 1979) and compute the RI variable according to a three-step procedure.

Firstly, for each case (RTA_{ir}^t) in any period we distinguish, within the network \mathcal{N} , those nodes exhibiting a value of $RTA_{ir}^t \geq 1$ from those having a $RTA_{ir}^t < 1$.

Secondly, we compute the number of times a node i is lying on the geodesic paths linking nodes j and k whose $RTA_{jr}^t \geq 1$.

Finally, we weighted each value by a constant value, $(N - 1) \times (N - 2) \times 2$ (i.e. 28560) to normalize each value for the European MST average. Formally:

$$RI_{ir}^t = \sum_{j \neq i}^{N-1} \frac{g_{jk}(n_i)}{(N - 1) \times (N - 2) \times 2} |RTA_{jr}^t, RTA_{kr}^t > 1, MST^t)$$

In synthesis, RI explains the technological specialization of a region, in a given IPC class, in terms of its strategic positioning within the MST. In other words, the technological specialization of a region in a given IPC class depends on the extent of potential combination of locally established specializations in unrelated technologies (i.e. on its bridging or gatekeeping role of that given IPS).

Exogenous Technological Change (ETC)

All previous variables are computed considering the structure of the MST and RTA with reference to the same period. However, in such a way we are unable to disentangle, for each region, RTA_{ir} , the effects played by its previous technological structure from those arising from exogenous changes in the European “technological frontier” which, by definition, are an exogenous phenomenon from the regional viewpoint. We attributed the original intuition of this concept to Michael Kalecki who wrote “The intensity of the technical progress of a society and its path of economic development is governed by the extent of such major exogenous innovation. A ceiling on the rate of growth of capital accumulation is determined by the level of adoption of the major technology within any particular economy” (Kalecki, 1954, p. 175).

For this reason, we compute ETC as a variable similar to LTC, but with a relevant difference: while in the LTC variable, the RTA and the MST are contemporaneous, in computing the ETC variable the MST is one period ahead ($t+1$) with respect to the value of the RTA (t). In this way, we are able to see whether past specialization of previously distant IPC sectors, which became technologically proximate in the subsequent period, played a role in determining the relative specialization in a specific IPC class in a given region. Being used in conjunction with LTC, which accounts for the effects driven by past MST, this variable can account for the effect of exogenous technological change in shaping a region’s technological specialization. Formally:

$$ETC_{ir}^t = \sum_{j=1}^{N-1} RTA_{jr}^t | (g_{ij} = 1, MST^{t+1})$$

where it can be noted that while ETC refers to time t , the MST on which it is built refers to time $t+1$.

In synthesis, ETC explains the technological specialization of a region, in a given IPC class, in terms of the current exogenous shifts and changes in the European MST⁶.

Proximity factors

In this study we assess the role of LTC, RI and ETC by also accounting for proximity factors along both the geographic and the technological dimensions.

The geographic matrix (W_{geo}) is computed as the inverse of the distance matrix between centroids of each region in the sample. The technological matrix (W_{tec}^t) is computed on the basis of socio-cognitive data. Each element of the W_{tec}^t matrix measures co-inventorships for couplets of regions. Different from the geographic matrix, it changes over time. Following Kelejian and Prucha (2010), both matrices are maximum eigenvalue normalized. The variables accounting for proximity in the

⁶ For robustness purposes, the variables LTC, RI and ETC are also constructed from the MST computed for each period from the CO-CL connectivity matrix.

geography and technological domains are then obtained by pre-multiplying the main explanatory variables (LTC, RI, ETC) by each of the normalized weight matrices (W_{geo}^n and $W_{tec}^{t,n}$).

5. Econometric models and methodological issues

We model the dependent variable, $SRTA_{ir}^t$, by means of ordered response models. According to Cameron and Trivedi (2009), ordered outcomes arise sequentially as a latent variable (*spec*) crosses increasingly higher thresholds as a function of explanatory variables (X) and controls:

$$spec_{ir}^t = \alpha_{ir} + X_{ir}^t \beta + controls + \varepsilon_{ir}^t$$

In our analysis, the X matrix contains the main explanatory variables (LTC, RI, ETC). In the case of a 5 ordered states model, we have

$SRTA_{ir}^t = k$ if $\gamma_{k-1} < spec_{ir}^t \leq \gamma_k$ with $k = 1 =$ Inactive, $2 =$ Very Unspecialised, ..., $K = 5 =$ Very Specialized.

$$P(SRTA_{ir}^t = k) = P(\gamma_{k-1} < spec_{ir}^t \leq \gamma_k) = P(\gamma_{k-1} < \alpha_{ir} + X_{ir}^t \beta + controls + \varepsilon_{ir}^t \leq \gamma_k) \quad (1)$$

$$P(SRTA_{ir}^t = k) = F[\gamma_k - (\alpha_{ir} + X_{ir}^t \beta + controls)] - F[\gamma_{k-1} - (\alpha_{ir} + X_{ir}^t \beta + controls)]$$

where F is the cumulative distribution function of the error term. We specify F as the cumulative logistic function, which yields the panel-ordered logit model.

Among the controls, we include the variables $W_{geo}^n X_{ir}^t$ and $W_{tec}^n X_{ir}^t$ to account for proximity factors along the geographic and the technological dimension, respectively. We also include dummies to account for the fact that the focal case can be the result of a forward or backward transition between states or a case of no-change with respect to the previous period. This way, we account for possible dynamics effects which could affect the role played by the main explanatory variables.

Time dummies and countries dummy are also included to account for macro shocks and national institutional features.

Finally, to attenuate the potential problem of endogeneity, which could arise because of possible simultaneity, all the explanatory variables are included in the model with a one-period lag. Given that such lag refers to the average over the previous five years, it is supposed to be sufficiently long to break the correlation between the error term and each of the regressors.

Following Wooldridge (2005, 2010), we estimate the panel logit-ordered models adopting the Correlated Random Effects (CRE) approach, which can be

seen as a unifying framework encompassing both the *fixed* (FE) and the *random* effects (RE) estimation approaches. The framework was first proposed by Mundlak (1978) and modified by Chamberlain (1980, 1982). It handles the correlation between the unobserved case effect, α_{ir} , and the time-varying regressors. More specifically, α_{ir} is modelled as a function of the time-averages ($\bar{Z}_{ir} \in \bar{X}_{ir}, \overline{W_{geo}^n X_{ir}}, \overline{W_{tec}^{t,n} X_{ir}}$) of all time-varying exogenous variables:

$$\alpha_{ir} = \mu + \bar{Z}_{ir}\xi_1 + v_{ir}$$

where v_{ir} has zero mean and is assumed to be uncorrelated with the regressors. Therefore, model (1) can be reformulated as:

$$P(SRTA_{ir}^t = k) = (\gamma_{k-1} < \mu + X_{ir}^t\beta + \bar{Z}_{ir}\xi_1 + controls + v_{ir} + \varepsilon_{ir}^t \leq \gamma_k) \quad (2)$$

The RE estimator can be used to consistently estimate all the coefficients. It yields the fixed effect estimates for β , while $\xi_1 = \hat{\beta}_B - \hat{\beta}_{FE}$, where $\hat{\beta}_B$ is the between estimator. Setting $\xi_j = 0$ ($j = 1, 2, 3$) results in inconsistent estimates of β , which is the problem of adopting the RE approach without accounting for the correlation between the individual effects and the regressors. On the other hand, if the FE estimator is adopted only the *within* variation is exploited while the *between* variation is discarded.⁷ The CRE approach is to be preferred as it yields consistent estimators while (as is the case for the RE estimator) accounting in an efficient way for both the within and the between variation.

6. Empirical analysis

The main results of the econometric analysis are reported in four tables, Tables 2-5. As mentioned above, our empirical estimation procedure is built along a bipartite perspective: a static and a dynamic one. Tables 2 and 3 display the results of the static perspective, whereas Tables 4 and 5 show the results of the dynamic perspective.

In the first two columns of Table 2, we present two different model specifications of the static perspective – without and with geographical and spatial spillover effects – of our baseline regression model. The explanatory variables included are obtained from the CO-SP European knowledge space. The third column replicates the second specification on a different European knowledge space obtained from the CO-CL matrix, for robustness checks. Average marginal effects

⁷ Bell and Jones (2015) argue that a variable that has a hierarchical structure can be decomposed in its between and within components, $x_{it} = x_i^B + x_{it}^W$, which can have different effects on the dependent variable.

for specification of the second column of Table 2 are reported in Table 3. Table 4 presents the results of the dynamic perspective for the three different groups as described in section 4.1, i.e. *stable*, *decreased* or *increased* specialization dynamics; Table 5 reports the average marginal effects for each model reported in Table 4.

TABLES 2-5 HERE

Turning to the results in Table 2, it is worth emphasizing that, thanks to the CRE approach, for each explanatory variable we are able to distinguish two types of effects, the within and the between effects. As the former refer to the effect due to a change in a given variable from one period to the following, this could be interpreted as short-run effects. On the contrary, the between effects, which are obtained by exploiting the cross-section variation in the data, could be interpreted as long-run effects. It is important to remark that the coefficients reported in Table 2 for the time averages of LTS, RI and ETC are the difference between the *between* and the *within* effects (net-between coefficients), as explained in the previous section.

All the models in Table 2 were estimated by including transition dummies to account for the fact that a case (region-IPC class observation) can be the result of a forward or a backward transition dynamics, the reference case is the no-change case. More specifically, we include three forward transition dummies (for 1 step, 2 steps and 3+4 steps ahead transitions) and three backward transition dummies (for 1 step, 2 steps and 3+4 steps backward transitions). Therefore, the effects of the main explanatory variables are not influenced by transition dynamics.

Focusing on the first two models of Table 2 all the coefficients of the explanatory variables exhibit the expected positive sign, this implies that when a given variable increases the probability of being in the *inactive* or *very unspecialised* state decreases, whereas the probability of being in the *unspecialised specialized*, *specialized* or *very specialized* state increases (see Table 3). When controls for geographical and technological spillovers arising from interregional flows of scientific and technological knowledge are included (second column), the magnitude of the coefficients remains stable, but the time average of the ETC variable turns out to be not significant. In general, the within coefficients are much lower when compared to the net-between coefficients. Focusing on the second model, LTC exhibits a coefficient of 0.0221, whereas its time average a coefficient of 0.2832; in the case of RI both coefficients are much lower, 0.0016 and 0.0177; for ETC the coefficients are 0.0033 and 0.0405. This finding indicates that cross-section variation plays a major role in explain the overall variation in the outcome variable. The coefficients of the transition dummies, all highly significant, exhibit the expected signs and have a very sizeable magnitude compared to the coefficients of the explanatory variables. This indicates the importance of accounting for the past states of the local technological cases analysed.⁸

⁸ As for the spatial and technological spillovers (not shown in detail in Table 2), the time average associated with LTC display positive and significant coefficients, while the

Results based on the Co-Classification approach to map the European knowledge space, and displayed in column 3 are overall comparable to the ones reported for model 2. One noticeable exception is that in model 3 RI is not significant, whereas the opposite is the case for the time average of ETC.

In Table 3, we report the average marginal effects for our preferred model, model 2 of Table 2. The within average marginal effects are lower in size with respect to the net-between ones. For all the variables, the negative effect on the probability of being in the *Inactive* state is larger in size with respect to the one exerted on the probability of being on the *Very unspecialised* state. For the other three states, the positive marginal effects exhibit a nonlinear behaviour, the larger effect is for the *Unspecialised* state, followed by the one for the *Very specialised* state, and the lowest positive effect is for the *Specialised* state. This indicates that the explanatory variables play a crucial role when regions are building their technological comparative advantage, i.e. when their RTA in a given IPC class is in the range from 0.5 to 1. A relevant effect is then played to reinforce specialization when $RTA > 1.5$. For all the explanatory variables, it is worth emphasizing that in absolute terms the largest effect is found for the *Inactive* ($RTA=0$) state. Being the effect negative, this indicates that all the variables included in our analysis play their major role in activating the process of specialization (RTA switches to positive values) at the local level. Comparing the effects across the three main variables, we find that the most effective one is LTC, followed by ETC. The smallest effects are associated with RI. This latter finding indicates how more complicated is for a region to gain and develop RTA in an IPC class when this is the result of the complex process of recombining in a successful way existing knowledge and ideas.

Overall, the results presented in Tables 2 and 3 provide sound evidence that the current technological specialisation of a region in a specific IPC class depends on the specialization of that region in contiguous IPC patent classes (LTC), on the crucial positioning (or betweenness) of a given IPC class as bridge a/o gatekeeper between other different technological specialization of the region (RI), and on the exogenous evolution of the European technological frontier (ETC).

In Table 4 we present the results for the three subsamples corresponding to the groups of *stable*, *increased* or *decreased* specialization dynamics. Average marginal effects are reported in Table 5. For the *stable* specialization dynamics, which represents almost 52% of the overall cases, the estimated coefficients are qualitatively similar to those obtained for the second model of Table 2: the within coefficients display a comparable magnitude, whereas the coefficients of the time average terms for LTC and RI are larger; the ETC time average is again not significant at conventional levels. For the *increased* specialization dynamics subsample (which accounts for 25% of the cases) we find that only the time average variable

coefficient for RI and ETI display often negative values. This result can be interpreted as evidence of the role played by tacit knowledge which can easily transferred across contiguous IPC classes but cannot be transferred over longer geographic and technological distances (as expressed by both weight matrices).

coefficients turn out to be significant. On the contrary, for the *decreased dynamics* subsample (23% of the cases) both within and between variations are relevant, although not all the coefficients are significant. More specifically, LTC exhibits a positive and significant coefficient only for its time average term, while both coefficients of RI are positive and significant. ETC does not seem to play any significant role (the within coefficient is significant at the 11%).

Focusing on the average marginal effects for the *stable* dynamic specialization group (Table 5), we find that when a given variable increases the probability of being in the first state (*inactive*) decreases, whereas the probability of being in all the other four states increases. In absolute terms and for all the variables, the largest effects are found for the *inactive* cases. The positive effects follow a nonlinear pattern, the largest effect is found for *unspecialized* cases, followed by *very unspecialised*, *very specialized* and *specialized* cases. As for the relative role played by the different variables, the most effective ones are the time averages of LTC and RI, followed by the within counterparts of LTC, ETC and RI.

A different picture emerges for the *increased* dynamic group, positive changes in the explanatory variables cause a decrease in the probability of being in the *very unspecialised* or in the *unspecialized* states. At the same time, such changes determine an increase in the probability of being in *specialized* or *very specialized* states. Only the average marginal effects of the time average variables are significant, pointing that only cross-section variation matters for positive transitions into higher states of technological comparative advantage. LTC is confirmed as the most effective variable, followed by ETC and RI. For all variables, the largest marginal effects are found for the *very specialized* cases.

The results for the *decreased* specialization dynamic group are mixed. For the RI variable, both the within and the net-between marginal effects are significant, for LTC only the marginal effect associated with the time average is significant, whereas ETC turns out to be not relevant. Again, time average variables' marginal effects are found to be larger than the within ones, with LTC being more effective than RI. The largest effects are found for the *inactive* state, showing the role of contiguous IPC classes in avoiding the total disappearance of a given characteristic in the spectrum of the technological specialization of a region. Taken all together, the results point to the fact that, in general, external forces act as re-enforcer of an internal process (of specialization or de-specialization). Such a process derives from the history of an IPC class in a given region and there is only a limited range of values (of the RTA specialization index) in which they may determine the “upward” or “downward” evolution of a region in that IPC class.

Overall, our analysis has provided stimulating insights on the determinants of technological specialization of the European regions and how their effects change over time when taking into proper account the role of geographic and technological connectivity among regions, as well as transition dynamics.

7. Conclusions

The technological and productive specialization of regions has always been a relevant issue both from a theoretical and empirical viewpoint. We contribute to the current debate by offering a novel empirical analysis focused on 198 most innovative European regions. Data on patents classified in 121 IPC sectors observed over the period 1986-2010 are used to map the European knowledge space based on conditional co-specialisations of regions in the same IPC Classes (as in Hidalgo et al., 2007). Thanks to this representation of the knowledge space, we investigate the evolution of the specialization process, measured in terms of the sector-region revealed technological advantage. The analysis is carried out within a novel approach, which we propose as a unifying framework that encompasses the three theoretical contributions of “localised technological change” (Atkinson and Stiglitz, 1969), “recombinant growth” (Weitzman, 1998) and exogenous innovation (Kalecki, 1954) within a SNA derived analytical framework.

Our empirical analysis, carried out by estimating panel-ordered logit spatial models, has provided convincing evidence of the role played by localised technological change, knowledge recombination, exogenous technological shifts, and spillovers arising from both geographic and technological regional proximity.

The main results show that regional technological specialization is mainly shaped by localised technological change and exogenous technological paradigm shifts, whereas processes of recombinant innovation contribute to a lower extent. Moreover, results also point out that, once accounting for spatial and technological spillovers and transition dynamics, it is the between-variation effect which plays a major role. When we split the sample to focus specifically on different specialization dynamics, results from the static analysis are confirmed but new evidence emerges, too. For stable cases, our independent variables exert their greater influence when a region has not developed any technological sector, followed by the case in which its RTA records values around 1. When the technological specialization is increasing over time the largest effects of independent variables is found for very specialised cases. Finally, when specialization decreases, the largest influence of independent variables is recorded with the lowest possible state, showing the role of contiguous IPC classes in avoiding the total disappearance of a given class in the spectrum of the technological specialization of a region.

Based on these results, future research may further deepen the understanding of the evolution of regional technological specialization by focusing on specific homogenous sub-groups of regions in order to investigate how specialization might have driven their current socio-economic performance and reinforced historical regional innovation divides.

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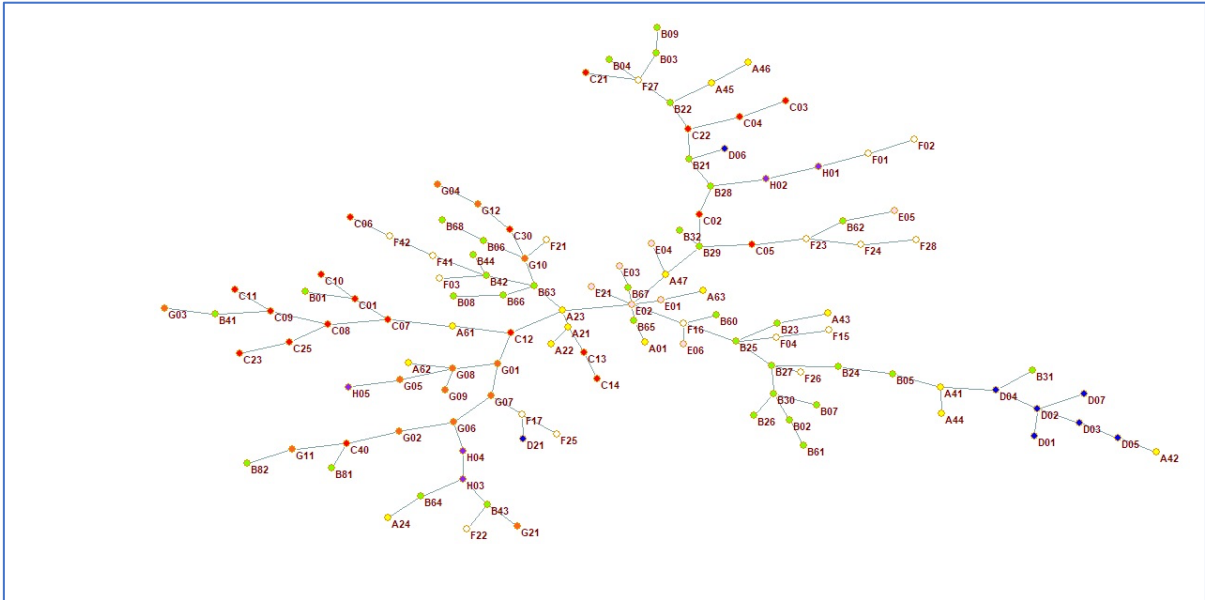
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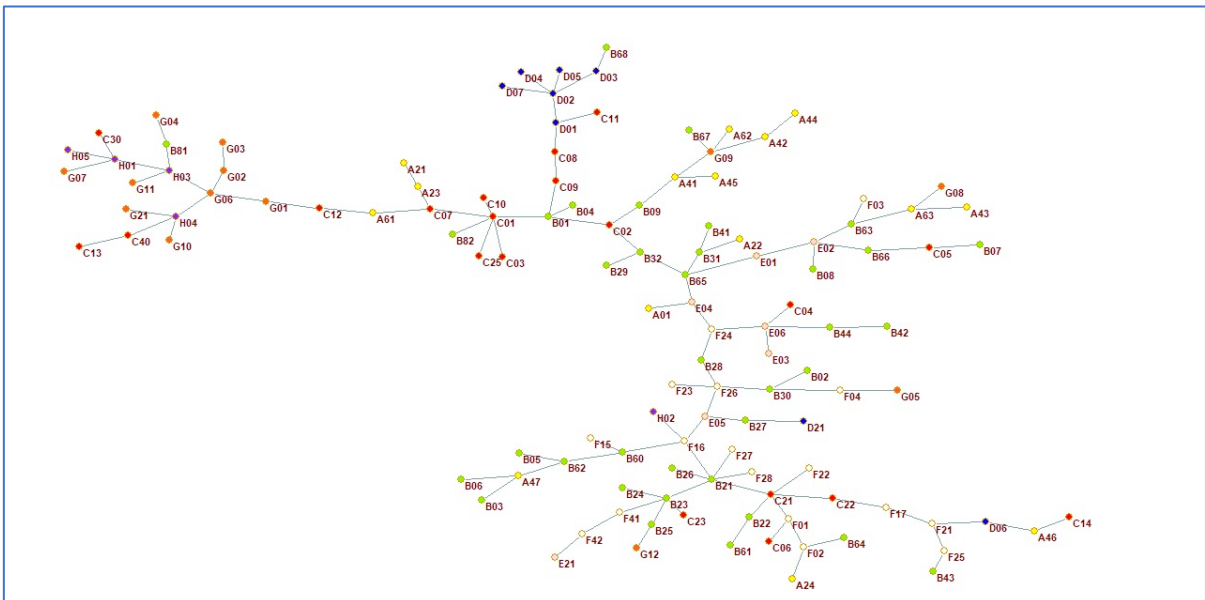
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FIGURES AND TABLES

Figure 1 – MST of the European knowledge space



T1: 1986-1990



T5: 2005-2010

Legend: A: Human Necessities (15 IPC classes at second hierarchy level); B: Performing operations; transporting (36); C: Chemistry; metallurgy (20); D: Textiles; paper (8); E: Fixed constructions (7); F: Mechanical engineering; lighting; heating; weapons; blasting (17); G: Physics (13); H: Electricity (5).

Figure 2 Distribution of SRTA among 5 specialisation states, all periods

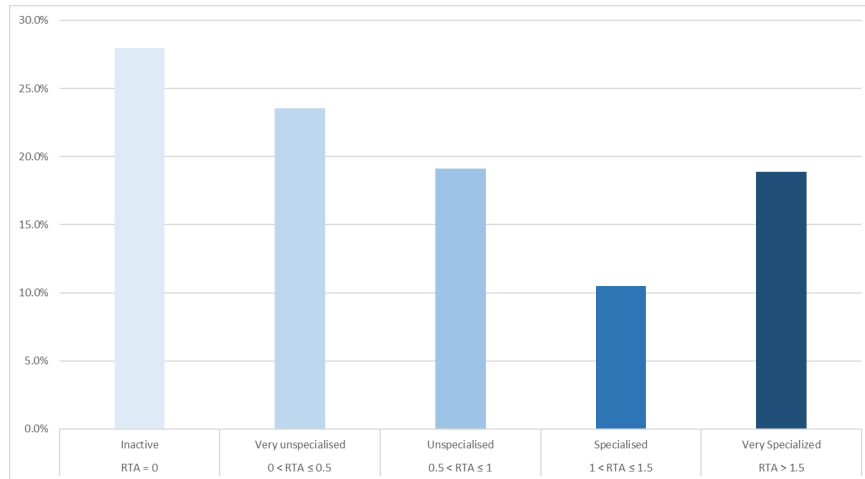


Figure 3: Distribution of RCA groups, all periods

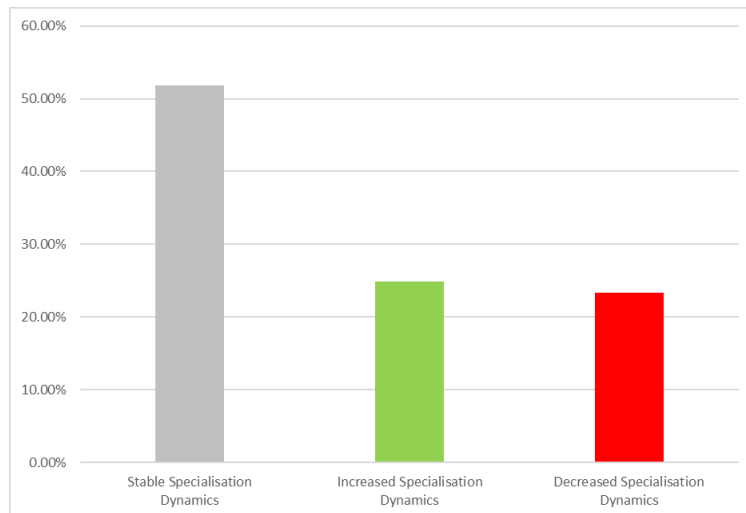


Table 1 - Distribution of RTA among the 5 static states and the 3 evolutive categories

		Panel A					
		<i>time t+1</i>					
		Absent	Traces	Not Specialized	Specialized	Very Specialized	Total
<i>time t</i>	Absent	19499	3515	1985	1170	3058	29227
	Traces	2483	11403	4853	1319	1133	21191
	Not Specialized	1421	4942	6604	2575	1913	17455
	Specialized	869	1452	2739	2494	2271	9825
	Very Specialized	2522	1249	2132	2521	9710	18134
	Total	26794	22561	18313	10079	18085	95832

		Panel B			
		1 step	2 steps	3 steps	4 steps
no change	49710				
forward change	23792	13214	5217	2303	3058
backward change	22330	12685	5005	2118	2522

Panel C

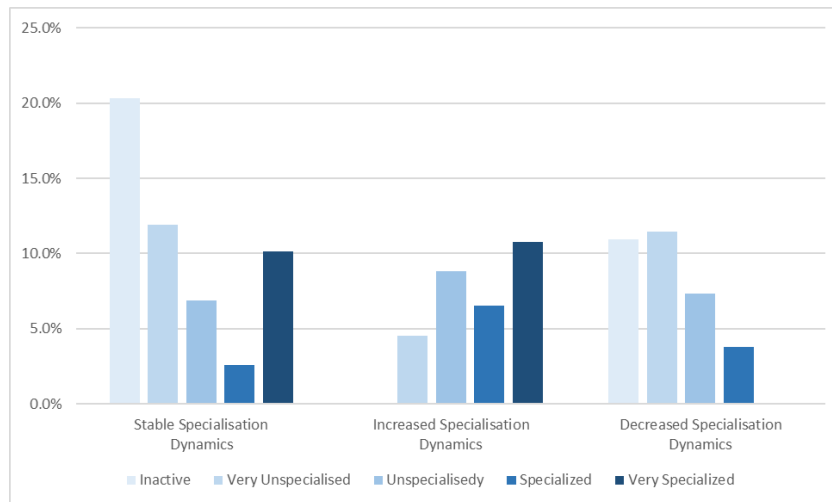


Table 2 - Correlated Random Effect Ordered Logit models for RTA states

	(1)	(2)	(3)
	Co-specialization		Co-classification
Localized technological change (LTC)	0.0229 *** (0.0057)	0.0221 *** (0.0057)	0.0187 *** (0.0048)
Recombinant Innovation (RI)	0.0005 *** (0.0001)	0.0016 *** (0.0002)	0.0005 *** (0.0002)
Exogenous technological change (ETC)	0.0029 ** (0.0012)	0.0033 *** (0.0013)	0.0010 (0.0015)
LTC - time average	0.3269 *** (0.0482)	0.2832 *** (0.0446)	0.1975 *** (0.0291)
RI - time average	0.0139 *** (0.0010)	0.0177 *** (0.0015)	0.0041 *** (0.0012)
ETC - time average	0.0524 * (0.0312)	0.0405 (0.0276)	0.0443 ** (0.0188)
<i>Dummies for transitions</i>			
1 step forward	1.9147 *** (0.0336)	1.9137 *** (0.0335)	1.9173 *** (0.0341)
2 steps forward	4.0409 *** (0.0789)	4.0554 *** (0.0793)	4.0530 *** (0.0785)
3 or 4 steps forward	8.6777 *** (0.1980)	8.7057 *** (0.1981)	8.7147 *** (0.1984)
1 step backward	-0.8578 *** (0.0552)	-0.8636 *** (0.0549)	-0.8433 *** (0.0553)
2 steps backward	-1.4111 *** (0.0854)	-1.4070 *** (0.0846)	-1.3675 *** (0.0855)
3 or 4 steps backward	-4.3674 *** (-0.1345)	-4.3284 *** (-0.1339)	-4.3307 *** (-0.1356)
<i>Geographic spillovers</i>	<i>no</i>	<i>yes</i>	<i>yes</i>
<i>Technological spillovers</i>	<i>no</i>	<i>yes</i>	<i>yes</i>
Pseudo Log-Likelihood	-101462	-100894	-102642

Time period: 1985-2010; observations refer to five-year averages. Number of observations 95832

All models include time dummies and country dummies

All explanatory variables are lagged one period

Standard errors are reported in parenthesis and clustered at the region level

Level of significance: *** 1%, ** 5%, * 10%

Table 3 - Average marginal effects computed for model 2 in Table 2

	<i>states</i>				
	Inactive	Very unspecialised	Unspecialised	Specialized	Very specialized
Model 2					
Localized technological change (LTC)	-0.002070 ***	-0.000524 **	0.001163 ***	0.000609 ***	0.000823 ***
Recombinant Innovation (RI)	-0.000146 ***	-0.000037 ***	0.000082 ***	0.000043 ***	0.000058 ***
Exogenous technological change (ETC)	-0.000311 **	-0.000079 **	0.000174 **	0.000091 **	0.000124 **
LTC - <i>time average</i>	-0.026516 ***	-0.006715 ***	0.014892 ***	0.007794 ***	0.010545 ***
RI - <i>time average</i>	-0.001659 ***	-0.000420 ***	0.000932 ***	0.000488 ***	0.000660 ***
ETC - <i>time average</i>	-0.003796	-0.000961	0.002132	0.001116	0.001510

Level of significance: *** 1%, ** 5%, * 10%

Table 4 - Correlated Random Effect Ordered Logit models - no-change and transition subsamples

	stable dynamic	emerged dynamic	declined dynamic
<i>RTA states</i>	1-5	2-5	1-4
Localized technological change (LTC)	0.0220 *** (0.0060)	-0.0156 (0.0114)	-0.0041 (0.0078)
Recombinant Innovation (RI)	0.0020 *** (0.0003)	0.0002 (0.0003)	0.0009 *** (0.0003)
Exogenous technological change (ETC)	0.0028 * (0.0017)	-0.0089 (0.0065)	0.0073 11% (0.0046)
LTC - time average	0.5236 *** (0.1201)	0.2287 *** (0.0219)	0.1207 *** (0.0203)
RI - time average	0.0343 *** (0.0044)	0.0022 *** (0.0005)	0.0083 *** (0.0007)
ETC - time average	0.0856 (0.0978)	0.0906 *** (0.0136)	0.0091 (0.0144)
<i>Geographic spillovers</i>	yes	yes	yes
<i>Technological spillovers</i>	yes	yes	yes
Pseudo Log-Likelihood	-46968.231	-30943.981	-27356.390
Observations	49710	23792	22330

Time period: 1985-2010; observations refer to five-year averages.

All models include period fixed effects and country dummies

All explanatory variables are lagged one period

Standard errors are reported in parenthesis and clustered at the region level

Level of significance: *** 1%, ** 5%, * 10%

Table 5 - Average marginal effects computed for Table 4 models

	<i>states</i>				
	Inactive	Very unspecialised	Unspecialised	Specialized	Very specialized
<i>Stable Specialisation Dynamics</i>					
Localized technological change (LTC)	-0.002154 ***	0.000658 **	0.000739 ***	0.000223 ***	0.000534 ***
Recombinant Innovation (RI)	-0.000196 ***	0.000060 **	0.000067 ***	0.000020 ***	0.000049 ***
Exogenous technological change (ETC)	-0.000272 *	0.000083	0.000093 *	0.000028 *	0.000067 *
LTC - <i>time average</i>	-0.051277 ***	0.015674 **	0.017584 ***	0.005302 ***	0.012718 ***
RI - <i>time average</i>	-0.003356 ***	0.001026 **	0.001151 ***	0.000347 ***	0.000832 ***
ETC - <i>time average</i>	-0.008386	0.002563	0.002876	0.000867	0.002080
<i>Increased Specialisation Dynamics</i>					
Localized technological change (LTC)		0.001703	0.001921	-0.000380	-0.003244
Recombinant Innovation (RI)		-0.000026	-0.000029	0.000006	0.000049
Exogenous technological change (ETC)		0.000971	0.001095	-0.000217	-0.001850
LTC - <i>time average</i>		-0.024914 ***	-0.028098 ***	0.005554 ***	0.047458 ***
RI - <i>time average</i>		-0.000245 ***	-0.000276 ***	0.000055 ***	0.000466 ***
ETC - <i>time average</i>		-0.009869 ***	-0.011131 ***	0.002200 ***	0.018800 ***
<i>Decreased Specialisation Dynamics</i>					
Localized technological change (LTC)	0.000797	-0.000085	-0.000433	-0.000279	
Recombinant Innovation (RI)	-0.000180 ***	0.000019 ***	0.000098 ***	0.000063 ***	
Exogenous technological change (ETC)	-0.001433	0.000152	0.000780	0.000502	
LTC - <i>time average</i>	-0.023631 ***	0.002506 ***	0.012853 ***	0.008272 ***	
RI - <i>time average</i>	-0.001626 ***	0.000172 ***	0.000884 ***	0.000569 ***	
ETC - <i>time average</i>	-0.001772	0.000188	0.000964	0.000620	

Level of significance: *** 1%, ** 5%, * 10%

APPENDIX

Table A1 - Countries and regions

Country	Regions	EU
Austria	9	
Belgium	11	
Switzerland	7	no
Germany	38	
Denmark	5	
Finland	5	
France	22	
Italy	21	
The Netherlands	12	
Norway	7	no
Spain	16	
Sweden	8	
United Kingdom	37	
Total	198	

Table A2 - QAP correlations for CO-SP MST

	T1	T2	T3	T4	T5
T1	1.000				
T2	0.220	1.000			
T3	0.220	0.246	1.000		
T4	0.170	0.187	0.246	1.000	
T5	0.170	0.170	0.212	0.212	1.000

Note: All coefficients are statistically significant at 1%.

Legend: time intervals are as follows: 1986-90 (T1), 1991-95 (T2), 1996-2000 (T3), 2001-05 (T4), 2006-10 (T5).

Table A3 - Variable definitions and descriptive statistics

Variable	Definition	mean	st. dev.	min	max
Revealed Technological Advantage (RTA)	Proportion of a region's patents in a given IPC class divided by the proportion of European patents in the same IPC class.*	1.083	3.013	0	261.5294
Localized technological change (LTC)	Sum of the RTA values of contiguous sectors in the European Maximum Spanning Tree	2.135	4.166	0	261.529
Recombinant Innovation (RI)	Betweenness centrality index in the European MST. The index is computed only for IPCs with RTA>1	40.943	97.593	0	1430.000
Exogenous technological change - (ETC)	Sum of the RTA values at time $t-1$ of contiguous sectors in the European Maximum Spanning Tree at time t che è quello della dipendente RTA t	2.180	5.688	0	760.428
$W_{geo}^n * LTC$	Geographical lags of the explanatory variables based on geographic proximity. The matrix W_{geo}^n is the inverse distance matrix, maximum eigenvalue normalized	1.280	1.001	0.077	26.174
$W_{geo}^n * RI$		25.842	54.304	0	918.000
$W_{geo}^n * ETC$		1.295	1.018	0.068	20.856
$W_{tec}^n * LTC$	Socio-cognitive lags of the explanatory variables based on technological proximity. The matrix W_{tec}^n is the regional co-inventorship matrix, maximum eigenvalue normalized	0.414	0.788	0	16.115
$W_{tec}^n * RI$		10.364	40.615	0	1190.000
$W_{tec}^n * ETC$		0.414	0.783	0.000	16.332

The primary source of the data is the European Patent Office (EPO).

* Europe refers to the countries included in the analysis: Austria, Belgium, Denmark, Germany, Spain, Finland, France, Italy, Netherlands, Sweden, United Kingdom, Switzerland and Norway
 Number of observations: 119790=121 IPC classes * 198 NUTS2 regions * 5 periods (time periods are 5 year-averages over the 1986-2010 years)

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